

# Hsu-Robbins theorem for the correlated sequences

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**CONFERENCE in honour of Prof. MAGDA PELIGRAD**  
**Paris, JUNE 2010**

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# Hsu-Robbins and Erdős

A famous result by Hsu and Robbins (1947) says that if  $X_1, X_2, \dots$  is a sequence of independent identically distributed random variables with zero mean and finite variance and

$$S_n := X_1 + \dots + X_n,$$

then

$$\sum_{n \geq 1} P(|S_n| > \varepsilon n) < \infty$$

for every  $\varepsilon > 0$ .

Note that, by the law of large numbers,

$$\frac{S_n}{n} \xrightarrow{n \rightarrow \infty} 0 = \mathbf{E}(X_1)$$

so

$$P(|S_n| > \varepsilon n) = P\left(\left|\frac{S_n}{n}\right| > \varepsilon\right) \xrightarrow{n \rightarrow \infty} 0$$

for every  $\varepsilon > 0$ .

So, the result of Hsu-Robbins says that if the variance of  $X_1$  is finite, this convergence is strong enough to ensure the summability of the series.

Later, Erdős (1949) showed that the converse implication also holds, namely if the series

$$\sum_{n \geq 1} P(|S_n| > \varepsilon n)$$

is finite for every  $\varepsilon > 0$  and  $X_1, X_2, \dots$  are independent and identically distributed, then  $\mathbf{E}X_1 = 0$  and  $\mathbf{E}X_1^2 < \infty$ .

Since then, many authors extended this result in several directions.

Spitzer's showed that

$$\sum_{n \geq 1} \frac{1}{n} P(|S_n| > \varepsilon n) < \infty$$

for every  $\varepsilon > 0$  if and only if  $\mathbf{E}X_1 = 0$  and  $\mathbf{E}|X_1| < \infty$ .

So, one introduces the factor  $\frac{1}{n}$  to "help" the convergence of the series and one needs a weaker conditions on the moments

Also, Spitzer's theorem has been the object of various generalizations and variants.

One of the problems related to the Hsu-Robbins' and Spitzer's theorems is to find the precise asymptotic as

$$\varepsilon \rightarrow 0$$

of the quantities

$$\sum_{n \geq 1} P(|S_n| > \varepsilon n)$$

and

$$\sum_{n \geq 1} \frac{1}{n} P(|S_n| > \varepsilon n)$$

.

Obviously, these sequences goes to  $\infty$  when  $\varepsilon \rightarrow 0$ .

Heyde (1975) showed that

$$\lim_{\varepsilon \rightarrow 0} \varepsilon^2 \sum_{n \geq 1} P(|S_n| > \varepsilon n) = \mathbf{E}X_1^2 \quad (1)$$

whenever  $\mathbf{E}X_1 = 0$  and  $\mathbf{E}X_1^2 < \infty$ . In the case when  $X$  is attracted to a stable distribution of exponent  $\alpha > 1$ , Spataru proved that

$$\lim_{\varepsilon \rightarrow 0} \frac{1}{-\log \varepsilon} \sum_{n \geq 1} \frac{1}{n} P(|S_n| > \varepsilon n) = \frac{\alpha}{\alpha - 1}. \quad (2)$$

It also holds for the Gaussian case : the limit is 2



# Variations of the fractional Brownian motion

Our purpose is to prove Hsu-Robbins and Spitzer's theorems for sequences of correlated random variables, related to the increments of fractional Brownian motion or to moving averages sequences

Recall that the fractional Brownian motion  $(B_t^H)_{t \in [0,1]}$  is a centered Gaussian process with covariance function

$R^H(t, s) = \mathbf{E}(B_t^H B_s^H) = \frac{1}{2}(t^{2H} + s^{2H} - |t - s|^{2H})$ . It can be also defined as the unique self-similar Gaussian process with stationary increments.

Concretely, in this paper we will study the behavior of the tail probabilities of the sequence

$$\begin{aligned} V_n &= \sum_{k=0}^{n-1} H_q \left( n^H \left( B_{\frac{k+1}{n}} - B_{\frac{k}{n}} \right) \right) \\ &\stackrel{(d)}{=} \sum_{k=0}^{n-1} H_q (B_{k+1} - B_k) \end{aligned} \quad (3)$$

where  $B$  is a fractional Brownian motion with Hurst parameter  $H \in (0, 1)$  (in the sequel we will omit the superscript  $H$  for  $B$ ) and  $H_q$  is the Hermite polynomial of degree  $q \geq 1$  given by

$$H_q(x) = (-1)^q e^{\frac{x^2}{2}} \frac{d^q}{dx^q} (e^{-\frac{x^2}{2}}).$$

If  $q = 1$  we have  $X_k = n^H \left( B_{\frac{k+1}{n}} - B_{\frac{k}{n}} \right)$  and

$$\mathbf{E}X_k X_l \neq 0$$

(unless  $H = \frac{1}{2}$ .)

If  $q = 2$  then

$$\begin{aligned} X_k &= H_2 \left( n^H \left( B_{\frac{k+1}{n}} - B_{\frac{k}{n}} \right) \right) \\ &= \left( n^H \left( B_{\frac{k+1}{n}} - B_{\frac{k}{n}} \right) \right)^2 - 1. \end{aligned}$$

In our case the variables are correlated. Indeed, for any  $k, l \geq 1$  we have

$$\mathbf{E} (H_q(B_{k+1} - B_k)H_q(B_{l+1} - B_l)) = \frac{1}{(q!)}\rho_H(k-l)^q$$

where the correlation function is

$$\rho_H(k) = \frac{1}{2} \left( (k+1)^{2H} + (k-1)^{2H} - 2k^{2H} \right)$$

which is not equal to zero unless  $H = \frac{1}{2}$  (which is the case of the standard Brownian motion).

# The convergence of the sequence $V_n$ .

Let  $q \geq 2$  an integer and let  $(B_t)_{t \geq 0}$  a fractional Brownian motion with Hurst parameter  $H \in (0, 1)$ . Then, with some explicit positive constants  $c_{1,q,H}$ ,  $c_{2,q,H}$  depending only on  $q$  and  $H$  we have

i. If  $0 < H < 1 - \frac{1}{2q}$  then

$$\frac{V_n}{c_{1,q,H}\sqrt{n}} \xrightarrow{n \rightarrow \infty} N(0, 1) \quad (4)$$

ii. If  $1 - \frac{1}{2q} < H < 1$  then

$$\frac{V_n}{c_{2,q,H}n^{1-q(1-H)}} \xrightarrow{n \rightarrow \infty} Z \quad (5)$$

where  $Z$  is a Hermite random variable (an iterated stochastic integral)

Example : for  $q = 2$  we have the quadratic variations of the fBm which converge as follows :

if

$$H < \frac{3}{4}$$

these variations converge (after normalization) to the normal law and

if

$$H > \frac{3}{4}$$

these variations converges (after normalization) to a non-Gaussian law (double stochastic integral, Rosenblatt)

Our purpose : prove precise asymptotics in Hsu-Robbins theorem for  $V_n$ , that is look to the quantities

$$\sum_{n \geq 1} P(V_n > \varepsilon n)$$

and

$$\sum_{n \geq 1} \frac{1}{n} P(V_n > \varepsilon n)$$

- no problems related to the existence of moments
- for every  $\varepsilon > 0$  the above series are convergent
- we will use chaos expansion and Malliavin calculus (the so -called Stein method)

# Multiple Wiener-Itô integrals

Let  $(W_t)_{t \in [0,1]}$  a standard Wiener process.

If  $f \in L^2([0, 1]^n)$  we define **the multiple Wiener integral of  $f$  with respect to  $W$**

Let  $f$  be a step function ( $f \in \mathcal{S}$ ), that means

$$f = \sum_{i_1, \dots, i_n} c_{i_1, \dots, i_n} 1_{A_{i_1} \times \dots \times A_{i_n}}$$

(here  $c_{i_1, \dots, i_n} = 0$  if two indices  $i_k$  and  $i_l$  are equal and the sets  $A_i \in \mathcal{B}([0, 1])$  are disjoint). We define for such a step function

$$I_n(f) = \sum_{i_1, \dots, i_n} c_{i_1, \dots, i_n} W(A_{i_1}) \dots W(A_{i_n})$$

where e.g.  $W([a, b]) = W_b - W_a$ .



We have that

- the application  $I_n$  is an isometry on  $\mathcal{S}$ , i.e.

$$\mathbf{E}(I_n(f)I_m(g)) = n! \langle f, g \rangle_{L^2([0,1]^n)} \text{ if } m = n$$

and

$$\mathbf{E}(I_n(f)I_m(g)) = 0 \text{ if } m \neq n$$

- the set  $\mathcal{S}$  is dense in  $L^2([0,1]^n)$

Therefore  $I_n$  can be extended to an isometry from  $L^2([0,1]^n)$  to  $L^2(\Omega)$ .

$I_n(f) = I_n(\tilde{f})$  where  $\tilde{f}$  is the symmetrization of  $f$

**Remark :**  $I_n$  can be viewed as an iterated stochastic Itô integral

$$I_n(f) = n! \int_0^1 \int_0^{t_n} \dots \int_0^{t_2} f(t_1, \dots, t_n) dW_{t_1} \dots dW_{t_n}$$

## Hermite random variable

The Hermite random variable of order  $q \geq 1$  that appears as limit in the above theorem is defined as

$$Z = d(q, H)I_q(L) \quad (6)$$

where the kernel  $L \in L^2([0, 1]^q)$  is given by

$$L(y_1, \dots, y_q) = \int_{y_1 \vee \dots \vee y_q}^1 \partial_1 K^H(u, y_1) \dots \partial_1 K^H(u, y_q) du.$$

The constant  $d(q, H)$  is a positive normalizing constant that guarantees that  $\mathbf{E}Z^2 = 1$  and  $K^H$  is the standard kernel of the fractional Brownian motion. We will not need the explicit expression of this kernel. Note that the case  $q = 1$  corresponds to the fractional Brownian motion and the case  $q = 2$  corresponds to the Rosenblatt process.

Let us denote, for every  $\varepsilon > 0$ ,

$$f_1(\varepsilon) = \sum_{n \geq 1} \frac{1}{n} P(V_n > \varepsilon n) = \sum_{n \geq 1} \frac{1}{n} P\left(Z_n^{(1)} > c_{1,q,H}^{-1} \varepsilon \sqrt{n}\right) \quad (7)$$

where

$$Z_n^{(1)} = \frac{V_n}{c_{1,q,H} \sqrt{n}}$$

while if  $1 - \frac{1}{2q} < H < 1$ , we are interested in

$$f_2(\varepsilon) = \sum_{n \geq 1} \frac{1}{n} P \left( V_n > \varepsilon n^{2-2q(1-H)} \right) = \sum_{n \geq 1} \frac{1}{n} P \left( Z_n^{(2)} > c_{2,q,H}^{-1} \varepsilon n^{1-q(1-H)} \right) \quad (8)$$

where

$$Z_n^{(2)} = \frac{V_n}{c_{2,q,H} n^{1-q(1-H)}}$$

It is natural to consider the tail probability of order  $n^{2-2q(1-H)}$  because the  $L^2$  norm of the sequence  $V_n$  is in this case of order  $n^{1-q(1-H)}$ .

We are interested to study the behavior of  $f_i(\varepsilon)$  ( $i = 1, 2$ ) as  $\varepsilon \rightarrow 0$ .

For a given random variable  $X$ , we set

$$\Phi_X(z) = 1 - P(X < z) + P(X < -z).$$

The first lemma gives the asymptotics of the functions  $f_i(\epsilon)$  as  $\epsilon \rightarrow 0$  when  $Z_n^{(i)}$  are replaced by their limits.

Consider  $c > 0$ .

i. Let  $Z^{(1)}$  be a standard normal random variable.

Then as

$$\frac{1}{-\log c\epsilon} \sum_{n \geq 1} \frac{1}{n} \Phi_{Z^{(1)}}(c\epsilon\sqrt{n}) \xrightarrow{\epsilon \rightarrow 0} 2.$$

ii. Let  $Z^{(2)}$  be a Hermite random variable of order  $q$  given by (6). Then, for any integer  $q \geq 1$

$$\frac{1}{-\log c\epsilon} \sum_{n \geq 1} \frac{1}{n} \Phi_{Z^{(2)}}(c\epsilon n^{1-q(1-H)}) \xrightarrow{\epsilon \rightarrow 0} \frac{1}{1-q(1-H)}.$$

Let  $q \geq 2$  and  $c > 0$ .

i. If  $H < 1 - \frac{1}{2q}$ , let  $Z^{(1)}$  be standard normal random variable.

Then it holds

$$\frac{1}{-\log c\varepsilon} \left[ \sum_{n \geq 1} \frac{1}{n} P \left( |Z_n^{(1)}| > c\varepsilon\sqrt{n} \right) - \sum_{n \geq 1} \frac{1}{n} P \left( |Z^{(1)}| > c\varepsilon\sqrt{n} \right) \right] \\ \xrightarrow{\varepsilon \rightarrow 0} 0.$$



ii. Let  $Z^{(2)}$  be a Hermite random variable of order  $q \geq 2$  and  $H > 1 - \frac{1}{2q}$ . Then

$$\frac{1}{-\log c\varepsilon} \left[ \sum_{n \geq 1} \frac{1}{n} P \left( |Z_n^{(2)}| > c\varepsilon n^{1-q(1-H)} \right) - \sum_{n \geq 1} \frac{1}{n} P \left( |Z^{(2)}| > c\varepsilon n^{1-q(1-H)} \right) \right] \rightarrow_{\varepsilon \rightarrow 0} 0.$$

# Idea of the proof

For the part i), it is based on Stein's method and Malliavin calculus ( $F$  is arbitrary,  $Z \sim N(0, 1)$ ) (Nourdin -Peccati)

$$\sup_{z \in \mathbb{R}} |P(F < z) - P(Z < z)| = \sup_{z \in \mathbb{R}} |\mathbf{E} (f'_z(F) - Ff_z(F))|$$

where  $f_z$  is the solution of the Stein's equation

$$1_{(-\infty, z)}(x) - P(Z < z) = f'(x) - xf(x), \quad x \in \mathbb{R}.$$

Since

$$\mathbf{E}Ff(F) = \mathbf{E}\delta D(-L)^{-1}Ff(F) = \mathbf{E}f'(F)\langle D(-L)^{-1}F, DF \rangle$$

we obtain

$$\sup_{z \in \mathbb{R}} |P(F < z) - P(Z < z)| \leq (\mathbf{E}(1 - \langle DF, D(-L)^{-1}F \rangle)^2)^{\frac{1}{2}}$$

$D$  is the Malliavin derivative,  $L$  the Ornstein-Uhlenbeck operator

To get a feeling

$$D_s W_t = 1_{[0,t]}(s)$$

$$D_s W_t^n = n W_t^{n-1} D_s W_t$$

$$D_s I_n(f) = I_{n-1}(f_n(\cdot, s))$$

$$(-L)^{-1} I_n(f) = \frac{1}{n} I_n(f)$$

It follows that

$$\sup_{x \in \mathbb{R}} \left| P \left( Z_n^{(1)} > x \right) - P \left( Z^{(1)} > x \right) \right|$$

$$\leq c \begin{cases} \frac{1}{\sqrt{n}}, & H \in (0, \frac{1}{2}] \\ n^{H-1}, & H \in [\frac{1}{2}, \frac{2q-3}{2q-2}) \\ n^{qH-q+\frac{1}{2}}, & H \in [\frac{2q-3}{2q-2}, 1 - \frac{1}{2q}). \end{cases}$$

and this implies that

$$\sum_{n \geq 1} \frac{1}{n} \sup_{x \in \mathbb{R}} \left| P \left( Z_n^{(1)} > x \right) - P \left( Z^{(1)} > x \right) \right|$$

$$\leq c \begin{cases} \sum_{n \geq 1} \frac{1}{n\sqrt{n}}, & H \in (0, \frac{1}{2}] \\ \sum_{n \geq 1} n^{H-2}, & H \in [\frac{1}{2}, \frac{2q-3}{2q-2}) \\ \sum_{n \geq 1} n^{qH-q-\frac{1}{2}}, & H \in [\frac{2q-3}{2q-2}, 1 - \frac{1}{2q}). \end{cases}$$

and the last sums are finite (for the last one we use  $H < 1 - \frac{1}{2q}$ ).

We state now the Spitzer's theorem for the variations of the fractional Brownian motion. Let  $f_1, f_2$  be given by

$$f_1(\varepsilon) = \sum_{n \geq 1} \frac{1}{n} P(V_n > \varepsilon n) = \sum_{n \geq 1} \frac{1}{n} P(Z_n^{(1)} > c_{1,q,H}^{-1} \varepsilon \sqrt{n}) \quad (9)$$

and

$$f_2(\varepsilon) = \sum_{n \geq 1} \frac{1}{n} P(V_n > \varepsilon n^{2-2q(1-H)}) = \sum_{n \geq 1} \frac{1}{n} P(Z_n^{(2)} > c_{2,q,H}^{-1} \varepsilon n^{1-q(1-H)}) \quad (10)$$

i. If  $0 < H < 1 - \frac{1}{2q}$  then

$$\lim_{\varepsilon \rightarrow 0} \frac{1}{\log(c_{1,H,q}^{-1} \varepsilon)} f_1(\varepsilon) = 2.$$

ii. If  $1 > H > 1 - \frac{1}{2q}$  then

$$\lim_{\varepsilon \rightarrow 0} \frac{1}{\log(c_{2,H,q}^{-1} \varepsilon)} f_2(\varepsilon) = \frac{1}{1 - q(1 - H)}.$$



for every  $\varepsilon > 0$

$$g_1(\varepsilon) = \sum_{n \geq 1} P(|V_n| > \varepsilon n) \quad (11)$$

if  $H < 1 - \frac{1}{2q}$  and by

$$g_2(\varepsilon) = \sum_{n \geq 1} P(|V_n| > \varepsilon n^{2-2q(1-H)}) \quad (12)$$

if  $H > 1 - \frac{1}{2q}$ . and we estimate the behavior of the functions  $g_i(\varepsilon)$  as  $\varepsilon \rightarrow 0$ . Note that we can write

$$g_1(\varepsilon) = \sum_{n \geq 1} P(|Z_n^{(1)}| > c_{1,q,H}^{-1} \varepsilon \sqrt{n})$$

$$g_2(\varepsilon) = \sum_{n \geq 1} P(|Z_n^{(2)}| > c_{2,q,H}^{-1} \varepsilon n^{1-q(1-H)})$$

We decompose it as : for  $H < 1 - \frac{1}{2q}$

$$g_1(\varepsilon) = \sum_{n \geq 1} P \left( |Z^{(1)}| > c_{1,q,H}^{-1} \varepsilon \sqrt{n} \right) \\ + \sum_{n \geq 1} \left[ P \left( |Z_n^{(1)}| > c_{1,q,H}^{-1} \varepsilon \sqrt{n} \right) - P \left( |Z^{(1)}| > c_{1,q,H}^{-1} \varepsilon \sqrt{n} \right) \right].$$

and for  $H > 1 - \frac{1}{2q}$

$$g_2(\varepsilon) \\ = \sum_{n \geq 1} P \left( |Z^{(2)}| > \varepsilon c_{2,q,H}^{-1} n^{1-q(1-H)} \right) + \\ \sum_{n \geq 1} \left[ P \left( |Z_n^{(2)}| > c_{2,q,H}^{-1} \varepsilon n^{1-q(1-H)} \right) - P \left( |Z^{(2)}| > c_{2,q,H}^{-1} \varepsilon n^{1-q(1-H)} \right) \right]$$

## Theorem

Let  $q \geq 2$ . Let  $Z^{(1)}$  be a standard normal random variable,  $Z^{(2)}$  a Hermite random variable of order  $q \geq 2$ . Then

i. If  $0 < H < 1 - \frac{1}{2q}$ , we have

$$(c_{1,q,H}^{-1} \varepsilon)^2 g_1(\varepsilon) \xrightarrow{\varepsilon \rightarrow 0} 1 = \mathbf{E} Z^{(1)}.$$

ii. If  $1 - \frac{1}{2q} < H < 1$  we have

$$(c_{2,q,H}^{-1} \varepsilon)^{\frac{1}{1-q(1-H)}} g_2(\varepsilon) \xrightarrow{\varepsilon \rightarrow 0} \mathbf{E} |Z^{(2)}|^{\frac{1}{1-q(1-H)}}.$$

Joint work with Solesne Bourguin (Paris 1)

we will consider long memory moving averages defined by

$$X_n = \sum_{i \geq 1} a_i \varepsilon_{n-i}, n \in \mathbb{Z}$$

where the innovations  $\varepsilon_i$  are centered i.i.d. random variables having at least finite second moments and the moving averages  $a_i$  are of the form  $a_i = i^{-\beta} L(i)$  with  $\beta \in (\frac{1}{2}, 1)$  and  $L$  slowly varying towards infinity. The covariance function  $\rho(m) = \mathbf{E}(X_0 X_m)$  behaves as  $c_\beta m^{-2\beta+1}$  when  $m \rightarrow \infty$  and consequently is not summable since  $\beta > \frac{1}{2}$ . Therefore  $X_n$  is usually called long-memory or “long-range dependence” moving average.

Let  $K$  be a deterministic function which has Hermite rank  $q$  and satisfies  $\mathbf{E}(K^2(X_n)) < \infty$  and define

$$S_N = \sum_{n=1}^N [K(X_n) - \mathbf{E}(K(X_n))].$$

Suppose that the  $\alpha_i$  are regularly varying with exponent  $-\beta$ ,  $\beta \in (1/2, 1)$  (i.e.  $\alpha_i = |i|^{-\beta} L(i)$  and that  $L(i)$  is slowly varying at  $\infty$ ). Then

i. If  $q < (2\beta - 1)^{-1}$ , then

$$h_{k,\beta}^{-1} N^{\beta q - \frac{q}{2} - 1} S_N \xrightarrow{N \rightarrow +\infty} Z^{(k)} \quad (13)$$

where  $Z^{(q)}$  is a Hermite random variable of order  $q$

ii. If  $q > (2\beta - 1)^{-1}$ , then

$$\frac{1}{\sigma_{k,\beta}\sqrt{N}} S_N \xrightarrow{N \rightarrow +\infty} \mathcal{N}(0, 1) \quad (14)$$

with  $\sigma_{k,\beta}$  a positive constant.

Wu (2006), H-C Ho and T. Hsing (1997), Peligrad and Utev (1997)

Take also the innovations  $\varepsilon$  to be the increments of the Wiener process

$$\varepsilon_i = W_{i+1} - W_i$$

Take  $K = H_q$  the Hermite polynomials

Note that  $X_n$  can also be written as

$$\begin{aligned}
 X_n &= \sum_{i=1}^{\infty} \alpha_i (W_{n-i} - W_{n-i-1}) = \sum_{i=1}^{\infty} \alpha_i I_1(\mathbf{1}_{[n-i-1, n-i]}) \\
 &= I_1 \left( \underbrace{\sum_{i=1}^{\infty} \alpha_i \mathbf{1}_{[n-i-1, n-i]}}_{f_n} \right) = I_1(f_n). \tag{15}
 \end{aligned}$$



As  $K = H_q$ ,  $S_N$  can be represented as

$$\begin{aligned}
 S_N &= \sum_{n=1}^N [H_q(I_1(f_n)) - \mathbf{E}(H_q(I_1(f_n)))] = \frac{1}{q!} \sum_{n=1}^N [I_q(f_n^{\otimes q}) - \mathbf{E}(I_q(f_n^{\otimes q}))] \\
 &= \frac{1}{q!} \sum_{n=1}^N I_q(f_n^{\otimes q}) = \frac{1}{q!} I_q\left(\sum_{n=1}^N f_n^{\otimes q}\right).
 \end{aligned}$$

In order to apply the same techniques, we need the speed of convergence of  $Z_N = cS_N/\sqrt{N}$  to the normal law, that means, we need to bound

$$\sup_{z \in \mathbb{R}} |\mathbf{P}(Z_N \leq z) - \mathbf{P}(Z \leq z)|$$

we will evaluate the quantity

$$\mathbf{E} \left( \left( 1 - q^{-1} \|DZ_N\|_{\mathcal{H}}^2 \right)^2 \right).$$

(this is the bound obtained via Malliavin calculus).

We have

$$D_t Z_N = D_t \left( \frac{1}{\sigma\sqrt{N}} \sum_{n=1}^N I_q (f_n^{\otimes q}) \right) = \frac{q}{\sigma\sqrt{N}} \sum_{n=1}^N I_{q-1} (f_n^{\otimes q-1}) f_n(t)$$

and

$$\|DZ_N\|_{\mathcal{H}}^2 = \frac{q^2}{\sigma^2 N} \sum_{k,l=1}^N I_{q-1} (f_k^{\otimes q-1}) I_{q-1} (f_l^{\otimes q-1}) \langle f_k, f_l \rangle_{\mathcal{H}} \quad (16)$$

The multiplication formula between multiple stochastic integrals gives us that

$$\begin{aligned}
 & I_{q-1} \left( f_k^{\otimes q-1} \right) I_{q-1} \left( f_l^{\otimes q-1} \right) \\
 = & \sum_{r=0}^{q-1} r! \binom{q-1}{r}^2 I_{2q-2-2r} \left( f_k^{\otimes q-1-r} \tilde{\otimes} f_l^{\otimes q-1-r} \right) \langle f_k, f_l \rangle_{\mathcal{H}}^r.
 \end{aligned}$$

By replacing in (16), we obtain

$$\begin{aligned}
 & \|DZ_N\|_{\mathcal{H}}^2 \\
 = & \frac{q^2}{\sigma^2 N} \sum_{r=0}^{q-1} r! \binom{q-1}{r}^2 \sum_{k,l=1}^N I_{2q-2-2r} \left( f_k^{\otimes q-1-r} \tilde{\otimes} f_l^{\otimes q-1-r} \right) \langle f_k, f_l \rangle_{\mathcal{H}}^r
 \end{aligned}$$

## Theorem

*Under the condition  $q > (2\beta - 1)^{-1}$ ,  $Z_N$  converges in law towards  $Z \sim \mathcal{N}(0, 1)$ . Moreover, there exists a constant  $C_\beta$ , depending uniquely on  $\beta$ , such that, for any  $N \geq 1$ ,*

$$\sup_{z \in \mathbb{R}} |\mathbf{P}(Z_N \leq z) - \mathbf{P}(Z \leq z)| \leq C_\beta \begin{cases} N^{\frac{q}{2} + \frac{1}{2} - q\beta} & \text{if } \beta \in \left( \frac{1}{2}, \frac{q}{2q-2} \right] \\ N^{\frac{1}{2} - \beta} & \text{if } \beta \in \left[ \frac{q}{2q-2}, 1 \right) \end{cases}$$

$$f_1(\varepsilon) = \sum_{N \geq 1} \frac{1}{N} P(|S_N| > \varepsilon N).$$

when  $q > \frac{1}{2\beta-1}$

$$\begin{aligned} f_1(\varepsilon) &= \sum_{N \geq 1} \frac{1}{N} P\left(\sigma^{-1} \frac{1}{\sqrt{N}} |S_N| > \frac{\varepsilon \sqrt{N}}{\sigma}\right) \\ &= \sum_{N \geq 1} \frac{1}{N} P\left(|Z| > \frac{\varepsilon \sqrt{N}}{\sigma}\right) \\ &\quad + \sum_{N \geq 1} \frac{1}{N} \left[ P\left(\sigma^{-1} \frac{1}{\sqrt{N}} |S_N| > \frac{\varepsilon \sqrt{N}}{\sigma}\right) - P\left(|Z| > \frac{\varepsilon \sqrt{N}}{\sigma}\right) \right] \end{aligned}$$

where  $Z$  denotes a standard normal random variable.

## Proposition

When  $q > \frac{1}{2\beta-1}$ ,

$$\lim_{\varepsilon \rightarrow 0} \frac{1}{-\log(\varepsilon)} f_1(\varepsilon) = 2$$

and when  $q < \frac{1}{2\beta-1}$  then

$$\lim_{\varepsilon \rightarrow 0} \frac{1}{-\log(\varepsilon)} f_2(\varepsilon) = \frac{1}{1 + \frac{q}{2} - \beta q}.$$

It is also possible to give Hsu-Robbins type results, meaning to find the asymptotic behavior as  $\varepsilon \rightarrow 0$  of

$$g_1(\varepsilon) = \sum_{N \geq 1} P(|S_N| > \varepsilon N)$$

when  $q > \frac{1}{2\beta-1}$